

# Exponential Bounds for Discrete Time, Conditionally Symmetric Martingales with Bounded Increments

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## A Simple Example to Motivate this Work

Let

- $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space.
- $d > 0$  and  $\gamma \in (0, 1]$  be fixed numbers.
- $\{U_k\}_{k \in \mathbb{N}}$  be i.i.d. random variables where

$$\mathbb{P}(U_k = +d) = \mathbb{P}(U_k = -d) = \frac{\gamma}{2}, \quad \mathbb{P}(U_k = 0) = 1 - \gamma, \quad \forall k \in \mathbb{N}.$$

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Consider the Markov process  $\{X_k\}_{k \geq 0}$  where  $X_0 = 0$  and

$$X_k = X_{k-1} + U_k, \quad \forall k \in \mathbb{N}.$$

## Large Deviation Analysis

From Cramér's theorem in  $\mathbb{R}$ , for every  $\alpha \geq \mathbb{E}[U_1] = 0$ ,

$$\lim_{n \rightarrow \infty} -\frac{1}{n} \ln \mathbb{P}(X_n \geq \alpha n) = I(\alpha)$$

where the rate function is given by

$$I(\alpha) = \sup_{t \geq 0} \{t\alpha - \ln \mathbb{E}[\exp(tU_1)]\}.$$

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Let  $\delta \triangleq \frac{\alpha}{d}$ . Calculation shows that

$$I(\alpha) = \delta x - \ln \left( 1 + \gamma [\cosh(x) - 1] \right)$$

$$x \triangleq \ln \left( \frac{\delta(1 - \gamma) + \sqrt{\delta^2(1 - \gamma)^2 + \gamma^2(1 - \delta^2)}}{\gamma(1 - \delta)} \right).$$

## Large Deviation Analysis (Cont.)

Lets examine the martingale approach in this simple case, where

$$X_k = \sum_{i=1}^k U_i, \quad \forall k \in \mathbb{N}, \quad X_0 = 0$$

with the natural filtration

$$\mathcal{F}_k = \sigma(U_1, \dots, U_k), \quad \forall k \in \mathbb{N}, \quad \mathcal{F}_0 = \{\emptyset, \Omega\}.$$

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In this case, the martingale has bounded increments, and for every  $k \in \mathbb{N}$

$$\begin{aligned} |X_k - X_{k-1}| &\leq d \\ \mathbb{E}[(X_k - X_{k-1})^2 | \mathcal{F}_{k-1}] &= \text{Var}(U_k) = \gamma d^2 \end{aligned}$$

almost surely.

## Large Deviation Analysis (Cont.)

Via the martingale approach, the following exponential inequality follows:

$$\mathbb{P}(X_n - X_0 \geq \alpha n) \leq \exp \left( -n D \left( \frac{\delta + \gamma}{1 + \gamma} \parallel \frac{\gamma}{1 + \gamma} \right) \right)$$

where  $\delta \triangleq \frac{\alpha}{d}$ , and

$$D(p||q) \triangleq p \ln \left( \frac{p}{q} \right) + (1 - p) \ln \left( \frac{1 - p}{1 - q} \right), \quad \forall p, q \in [0, 1]$$

is the divergence (relative entropy) between the probability distributions  $(p, 1 - p)$  and  $(q, 1 - q)$ . If  $\delta > 1$ , then the above probability is zero.

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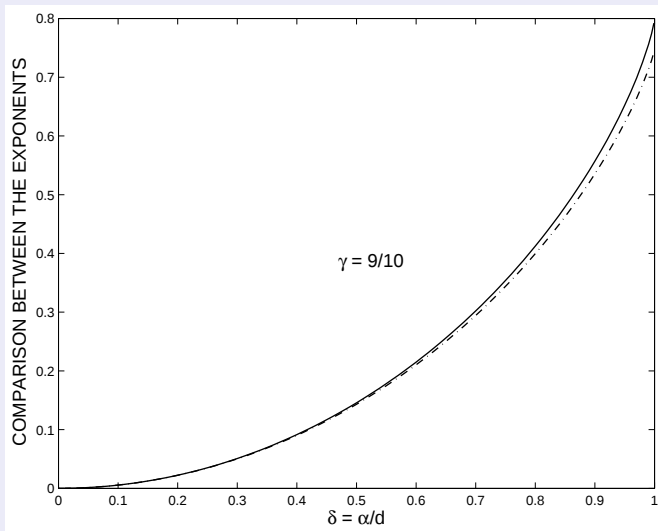
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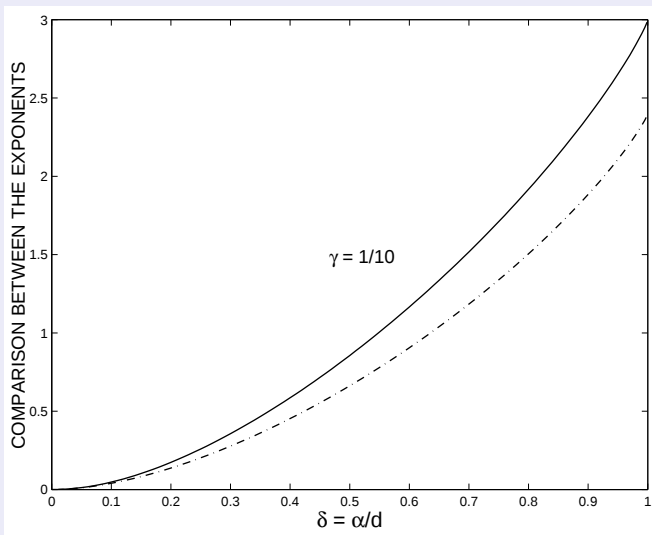
is the divergence (relative entropy) between the probability distributions  $(p, 1 - p)$  and  $(q, 1 - q)$ . If  $\delta > 1$ , then the above probability is zero.

This exponential bound is not tight (unless  $\gamma = 1$ ), and the gap to the exact asymptotic exponent increases as the value of  $\gamma \in (0, 1]$  is decreased.

## Comparison of the Exact Exponent and Its Lower Bound



## Comparison of the Exact Exponent and Its Lower Bound (Cont.)



## Questions:

- Why the lower bound on the exponent is not asymptotically tight ?
- Why this gap increases as the value of  $\gamma$  is decreased ?
- How this situation can be improved via the martingale approach ?

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This exponential bound relies on Bennett's inequality: Since  $U_k \leq d$  and  $\text{Var}(U_k) = \gamma d^2$ , then

$$\mathbb{E} [\exp(tU_k)] \leq \frac{\gamma \exp(td) + \exp(-\gamma td)}{1 + \gamma}, \quad \forall t \geq 0.$$

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The probability distribution that achieves Bennett's inequality with equality is asymmetric (unless  $\gamma = 1$ ), and is equal to

$$\mathbb{P}(\tilde{U}_k = d) = \frac{\gamma}{1 + \gamma}, \quad \mathbb{P}(\tilde{U}_k = -\gamma d) = \frac{1}{1 + \gamma}.$$

By reducing the value of  $\gamma \in (0, 1]$ , the above asymmetry grows. Indeed, this enlarges the gap to the exact asymptotic exponent.

## Interim Conclusion

An improvement is likely to be obtained by a refinement of Bennett's inequality when  $X_k$  is conditionally symmetric given  $\mathcal{F}_{k-1}$ .

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## Definition: Conditionally Symmetric Martingales

Let  $\{X_k, \mathcal{F}_k\}_{k \in \mathbb{N}_0}$ , where  $\mathbb{N}_0 \triangleq \mathbb{N} \cup \{0\}$ , be a discrete-time and real-valued martingale, and let  $\xi_k \triangleq X_k - X_{k-1}$  for every  $k \in \mathbb{N}$ . Then  $\{X_k, \mathcal{F}_k\}_{k \in \mathbb{N}_0}$  is a *conditionally symmetric martingale* if, conditioned on  $\mathcal{F}_{k-1}$ , the RV  $\xi_k$  is symmetrically distributed around zero.

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## Definition: Conditionally Symmetric Sub/ Supermartingales

Let  $\{X_k, \mathcal{F}_k\}_{k \in \mathbb{N}_0}$  be a discrete-time real-valued sub or supermartingale, and let  $\eta_k \triangleq X_k - \mathbb{E}[X_k | \mathcal{F}_{k-1}]$  for every  $k \in \mathbb{N}$ . Then it is conditionally symmetric if, conditioned on  $\mathcal{F}_{k-1}$ , the RV  $\eta_k$  is symmetrically distributed around zero.

## Construction of Conditionally Symmetric Martingales

Example 1: Let

- $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space
- $\{U_k\}_{k \in \mathbb{N}} \subseteq L^1(\Omega, \mathcal{F}, \mathbb{P})$  be independent random variables that are symmetrically distributed around zero
- $\{\mathcal{F}_k\}_{k \geq 0}$  be the natural filtration where  $\mathcal{F}_0 = \{\emptyset, \Omega\}$  and  $\mathcal{F}_k = \sigma(U_1, \dots, U_k)$ ,  $\forall k \in \mathbb{N}$ .
- For  $k \in \mathbb{N}$ , let  $A_k \in L^\infty(\Omega, \mathcal{F}_{k-1}, \mathbb{P})$  be an  $\mathcal{F}_{k-1}$ -measurable random variable with a finite essential supremum.

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Define a new sequence of random variables in  $L^1(\Omega, \mathcal{F}, \mathbb{P})$  where

$$X_n = \sum_{k=1}^n A_k U_k, \quad \forall n \in \mathbb{N}$$

and  $X_0 = 0$ . Then,  $\{X_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  is a conditionally symmetric martingale.

## Construction of Conditionally Symmetric Martingales

Example 2: Let

- $\{X_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  be a conditionally symmetric martingale
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Define  $Y_0 = 0$  and  $Y_n = \sum_{k=1}^n A_k(X_k - X_{k-1})$ ,  $\forall n \in \mathbb{N}$ .  
Then,  $\{Y_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  is a conditionally symmetric martingale.

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Then,  $\{Y_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  is a conditionally symmetric martingale.

Example 3: A sampled Brownian motion is a discrete time, conditionally symmetric martingale with un-bounded increments.

## Goal and Related Literature

- Our goal is to demonstrate how the assumption of the conditional symmetry improves existing exponential inequalities for discrete-time real-valued martingales with bounded increments.
- Earlier results, serving as motivation, appear in the works by
  - ① V. H. de la Pena (Annals of Probability 1999, 2004)
  - ② K. Dzhaparidze and J. H. van Zanten (SPA, 2001).
- The new exponential bounds in this work are also extended to conditionally symmetric sub or supermartingales.
- Additional results addressing weak-type inequalities, maximal inequalities and ratio inequalities for conditionally symmetric martingales were derived by
  - ① A. Oskowski (2010)
  - ② G. Wang (Proceedings of the AMS, 1991).

# Exponential Inequalities for Conditionally Symmetric Martingales

## Lemma

*Let  $X$  be a real-valued RV with a symmetric distribution around zero, a support  $[-d, d]$ , and assume  $\text{Var}(X) \leq \gamma d^2$  for some  $d > 0$  and  $\gamma \in [0, 1]$ . Let  $h$  be a real-valued convex function, and assume that  $h(d^2) \geq h(0)$ . Then,*

$$\mathbb{E}[h(X^2)] \leq (1 - \gamma)h(0) + \gamma h(d^2)$$

*with equality if  $\mathbb{P}(X = \pm d) = \frac{\gamma}{2}$ ,  $\mathbb{P}(X = 0) = 1 - \gamma$ .*

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## Proof

- $h$  convex,  $X \in [-d, d]$  a.s.  $\Rightarrow h(X^2) \leq h(0) + \left(\frac{X}{d}\right)^2 (h(d^2) - h(0))$ .
- Taking expectations on both sides gives the inequality, with an equality for the above symmetric distribution.

## Exp. Inequalities for Conditionally Symmetric Martingales (Cont.)

## Corollary

*Let  $X$  be a real-valued RV with a symmetric distribution around zero, a support  $[-d, d]$ , and assume  $\text{Var}(X) \leq \gamma d^2$  for some  $d > 0$  and  $\gamma \in [0, 1]$ . Then,*

$$\mathbb{E}[\exp(\lambda X)] \leq 1 + \gamma [\cosh(\lambda d) - 1], \quad \forall \lambda \in \mathbb{R}$$

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## Proof

- Symmetric distribution of  $X \Rightarrow \mathbb{E}[\exp(\lambda X)] = \mathbb{E}[\cosh(\lambda X)]$ .
- The corollary follows from the lemma since, for every  $x \in \mathbb{R}$ ,  $\cosh(\lambda x) = h(x^2)$  where  $h(x) \triangleq \sum_{n=0}^{\infty} \frac{\lambda^{2n} |x|^{2n}}{(2n)!}$  is a convex function, and  $h(d^2) = \cosh(\lambda d) \geq 1 = h(0)$ .

## Theorem 1

Let  $\{X_k, \mathcal{F}_k\}_{k \in \mathbb{N}_0}$  be a discrete-time real-valued and conditionally symmetric martingale. Assume that, for some fixed numbers  $d, \sigma > 0$ , the following two requirements are satisfied a.s.

$$|X_k - X_{k-1}| \leq d, \quad \text{Var}(X_k | \mathcal{F}_{k-1}) = \mathbb{E}[(X_k - X_{k-1})^2 | \mathcal{F}_{k-1}] \leq \sigma^2$$

for every  $k \in \mathbb{N}$ .

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for every  $k \in \mathbb{N}$ . Then, for every  $\alpha \geq 0$  and  $n \in \mathbb{N}$ ,

$$\mathbb{P} \left( \max_{1 \leq k \leq n} |X_k - X_0| \geq \alpha n \right) \leq 2 \exp(-nE(\gamma, \delta))$$

where

$$\gamma \triangleq \frac{\sigma^2}{d^2}, \quad \delta \triangleq \frac{\alpha}{d}$$

and for  $\gamma \in (0, 1]$  and  $\delta \in [0, 1)$ , the exponent  $E(\gamma, \delta)$  is given as follows:

## Theorem 1 (Cont.)

$$E(\gamma, \delta) \triangleq \delta x - \ln\left(1 + \gamma[\cosh(x) - 1]\right)$$

$$x \triangleq \ln\left(\frac{\delta(1 - \gamma) + \sqrt{\delta^2(1 - \gamma)^2 + \gamma^2(1 - \delta^2)}}{\gamma(1 - \delta)}\right).$$

If  $\delta > 1$ , then the probability is zero (so  $E(\gamma, \delta) \triangleq +\infty$ ), and  $E(\gamma, 1) = \ln\left(\frac{2}{\gamma}\right)$ .

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## Asymptotic Optimality of the Exponent

The example we studied earlier shows that the exponent of the bound in Theorem 1 is asymptotically optimal. That is, there exists a conditionally symmetric martingale, satisfying the conditions in this theorem, that attains the exponent  $E(\gamma, \delta)$  in the limit where  $n \rightarrow \infty$ .

Theorem 1 should be compared to the following existing bound which does not require the conditional symmetry property.

## Theorem 2

[McDiarmid 1989, book of Dembo & Zeitouni (Corollary 2.4.7)]

Let  $\{X_k, \mathcal{F}_k\}_{k \in \mathbb{N}_0}$  be a discrete-time real-valued martingale with bounded jumps. Assume that the two conditions on the bounded increments and conditional variance from Theorem 1 are satisfied a.s. for every  $k \in \mathbb{N}$ .

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$$\mathbb{P} \left( \max_{1 \leq k \leq n} |X_k - X_0| \geq \alpha n \right) \leq 2 \exp \left( -n D \left( \frac{\delta + \gamma}{1 + \gamma} \parallel \frac{\gamma}{1 + \gamma} \right) \right)$$

where  $\gamma$  and  $\delta$  are introduced in Theorem 1, and

$$D(p \parallel q) \triangleq p \ln \left( \frac{p}{q} \right) + (1 - p) \ln \left( \frac{1 - p}{1 - q} \right), \quad \forall p, q \in [0, 1].$$

If  $\delta > 1$ , then the probability is zero.

## Proof Technique for Theorems 1 and 2

- Define  $X_n - X_0 = \sum_{k=1}^n \xi_k$  where  $\xi_k \triangleq X_k - X_{k-1}$ .

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- Exponentiation and the maximal inequality for submartingales give that, for every  $\alpha \geq 0$ ,

$$\mathbb{P}\left(\max_{1 \leq k \leq n} (X_k - X_0) \geq n\alpha\right) \leq e^{-n\alpha t} \mathbb{E}\left[\exp\left(t \sum_{k=1}^n \xi_k\right)\right], \quad \forall t \geq 0.$$

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- Since  $\{\mathcal{F}_i\}$  forms a filtration, then for every  $t \geq 0$

$$\mathbb{E}\left[\exp\left(t \sum_{k=1}^n \xi_k\right)\right] = \mathbb{E}\left[\exp\left(t \sum_{k=1}^{n-1} \xi_k\right) \mathbb{E}[\exp(t\xi_n) | \mathcal{F}_{n-1}]\right].$$

## Proof Technique for Theorems 1 and 2 (Cont.)

- For proving Theorem 2, the use of Bennett's inequality gives

$$\mathbb{E} [\exp(t\xi_k) \mid \mathcal{F}_{k-1}] \leq \frac{\gamma \exp(td) + \exp(-\gamma td)}{1 + \gamma}.$$

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- For the refinement in Theorem 1 for conditionally symmetric martingales with bounded increments, use of Lemma 1 gives

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- In both cases, one gets exponential bounds with the free parameter  $t$ .
- Optimization over  $t \geq 0$  & union bound to get two-sided inequalities.
- The asymptotic optimality of the exponent in Theorem 1 follows from the simple example that was shown earlier.

### Theorem 3

Let  $\{X_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  be a discrete-time real-valued and conditionally symmetric martingale. Assume that there exists a fixed number  $d > 0$  such that  $\xi_k \triangleq X_k - X_{k-1} \leq d$  a.s. for every  $k \in \mathbb{N}$ . Let

$$Q_n \triangleq \sum_{k=1}^n \mathbb{E}[\xi_k^2 \mid \mathcal{F}_{k-1}]$$

with  $Q_0 \triangleq 0$ , be the predictable quadratic variation of the martingale up to time  $n$ .

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with  $Q_0 \triangleq 0$ , be the predictable quadratic variation of the martingale up to time  $n$ . Then, for every  $z, r > 0$ ,

$$\mathbb{P} \left( \max_{1 \leq k \leq n} (X_k - X_0) \geq z, Q_n \leq r \text{ for some } n \in \mathbb{N} \right) \leq \exp \left( -\frac{z^2}{2r} \cdot C \left( \frac{zd}{r} \right) \right)$$

where

$$C(u) \triangleq \frac{2[u \sinh^{-1}(u) - \sqrt{1+u^2} + 1]}{u^2}, \quad \forall u > 0.$$

Theorem 3 refines the following result that was stated without the requirement for the conditional symmetry of the martingale.

### Theorem 4 (Freedman 1975, Dembo & Zeitouni's book (Ex. 2.4.21))

Let  $\{X_n, \mathcal{F}_n\}_{n \in \mathbb{N}_0}$  be a discrete-time real-valued martingale. Assume that there exists a fixed number  $d > 0$  such that  $\xi_k \triangleq X_k - X_{k-1} \leq d$  a.s. for every  $k \in \mathbb{N}$ . Then, for every  $z, r > 0$ ,

$$\mathbb{P} \left( \max_{1 \leq k \leq n} (X_k - X_0) \geq z, Q_n \leq r \text{ for some } n \in \mathbb{N} \right) \leq \exp \left( -\frac{z^2}{2r} \cdot B \left( \frac{zd}{r} \right) \right)$$

where

$$B(u) \triangleq \frac{2[(1+u) \ln(1+u) - u]}{u^2}, \quad \forall u > 0.$$

## Related Refs. to this Work & Conditionally Symmetric Martingales

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I. Sason, "Tightened exponential bounds for discrete-time, conditionally symmetric martingales with bounded jumps," July 2012. Available at <http://arxiv.org/abs/1201.0533>.