

On Csiszár's f -Divergences and Informativities with Applications

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 - ▶ I-divergences (relative entropies);
 - ▶ χ^2 -divergence;
 - ▶ squared Hellinger distance;
 - ▶ total variation distance;
 - ▶ DeGroot statistical information;
 - ▶ etc.

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 - ▶ DeGroot statistical information;
 - ▶ etc.
- f -divergences satisfy the data processing inequality.

f -Informativities

f -Informativities (Csiszár, 1972) form a generalization of the mutual information:

- KL divergence \implies Shannon's Mutual Information;
- In general, f -divergence \implies f -informativity.

The Origins

- I. Csiszár, "Eine Informationstheoretische Ungleichung und ihre Anwendung auf den Beweis der Ergodizität von Markhoffschen Ketten," *Publ. Math. Inst. Hungar. Acad. Sci.*, vol. 8, pp. 85–108, Jan. 1963.
- I. Csiszár, "A note on Jensen's inequality," *Studia Scientiarum Mathematicarum Hungarica*, vol. 1, pp. 185–188, 1966.
- I. Csiszár, "Information-type measures of difference of probability distributions and indirect observations," *Studia Scientiarum Mathematicarum Hungarica*, vol. 2, pp. 299–318, Jan. 1967.
- I. Csiszár, "On topological properties of f -divergences," *Studia Scientiarum Mathematicarum Hungarica*, vol. 2, pp. 329–339, Jan. 1967.
- I. Csiszár, "A class of measures of informativity of observation channels," *Periodica Mathematicarum Hungarica*, vol. 2, pp. 191–213, Mar. 1972.
- S. M. Ali and S. D. Silvey, "A general class of coefficients of divergence of one distribution from another," *Journal of the Royal Statistics Society*, series B, vol. 28, no. 1, pp. 131–142, Jan. 1966.

Scope of this talk

Properties, and applications of f -divergences and f -informativities.

Notation

- \mathcal{C} denotes the set of convex functions $f: (0, \infty) \mapsto \mathbb{R}$ with $f(1) = 0$;
- P and Q are probability measures;
- $P, Q \ll \mu$ (e.g., $\mu = \frac{1}{2}(P + Q)$), and $p := \frac{dP}{d\mu}$, $q := \frac{dQ}{d\mu}$.

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 f -Divergence: Definition

The f -divergence from P to Q is given, independently of μ , by

$$D_f(P\|Q) := \int q f\left(\frac{p}{q}\right) d\mu \quad (1)$$

with the convention that

$$f(0) := \lim_{t \downarrow 0} f(t), \quad (2)$$

$$0f\left(\frac{0}{0}\right) := 0, \quad 0f\left(\frac{a}{0}\right) := \lim_{t \downarrow 0} tf\left(\frac{a}{t}\right) = a \lim_{u \rightarrow \infty} \frac{f(u)}{u}, \quad a > 0. \quad (3)$$

f-divergences: Examples

- Relative entropy

$$f(t) = t \log t, \quad t > 0 \implies D_f(P\|Q) = D(P\|Q), \quad (4)$$

$$f(t) = -\log t, \quad t > 0 \implies D_f(P\|Q) = D(Q\|P). \quad (5)$$

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- Total variation (TV) distance

$$f(t) = |t - 1|, \quad t \geq 0 \quad (6)$$

$$\implies D_f(P\|Q) = |P - Q| := \int \left| \frac{dP}{d\mu} - \frac{dQ}{d\mu} \right| d\mu, \quad P, Q \ll \mu. \quad (7)$$

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- Power divergence of order $\alpha \in (0, 1) \cup (1, \infty)$:

$$f_\alpha(t) = \frac{t^\alpha - \alpha(t - 1) - 1}{\alpha(\alpha - 1)}, \quad t \geq 0 \quad (8)$$

$$\implies \mathcal{I}_\alpha(P\|Q) := D_{f_\alpha}(P\|Q) := \frac{1}{\alpha(\alpha-1)} \left(\int \left(\frac{dP}{d\mu} \right)^\alpha \left(\frac{dQ}{d\mu} \right)^{1-\alpha} d\mu - 1 \right).$$

f-divergences: Examples (cont.)

- χ^2 -divergence:

$$\chi^2(P\|Q) := \int \frac{(p - q)^2}{q} d\mu = \frac{1}{2} \mathcal{I}_2(P\|Q). \quad (9)$$

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- **Relative entropies:** continuous extension at $\alpha = 0$ and $\alpha = 1$ yield

$$\mathcal{I}_1(P\|Q) = \frac{1}{\log e} D(P\|Q), \quad \mathcal{I}_0(P\|Q) = \frac{1}{\log e} D(Q\|P). \quad (10)$$

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- Squared Hellinger distance:

$$\mathcal{H}^2(P\|Q) := \frac{1}{2} \int (\sqrt{p} - \sqrt{q})^2 d\mu = 1 - \int \sqrt{pq} d\mu = \frac{1}{4} \mathcal{I}_{\frac{1}{2}}(P\|Q). \quad (11)$$

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- Rényi divergence of order $\alpha \in (0, 1) \cup (1, \infty)$:

$$D_\alpha(P\|Q) = \frac{1}{\alpha - 1} \log(1 + \alpha(\alpha - 1) \mathcal{I}_\alpha(P\|Q)). \quad (12)$$

Measures of Dependence (Rényi 1959, Csiszár 1967)

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- Csiszár suggested using f -divergences as dependence measures:
 $D_f(P_{XY} \| P_X \times P_Y)$ fulfills the postulates by Rényi if $f \in \mathcal{C}$ is strictly convex at 1, and $\lim_{t \rightarrow \infty} \frac{f(t)}{t} = +\infty$.

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▶ **Mutual information:**

$$f(t) = t \log t, (t > 0) \implies D_f(P_{XY} \| P_X \times P_Y) = I(X; Y). \quad (13)$$

▶ **Mean square contingency:** $f(t) = (t - 1)^2, (t \geq 0)$

$$\implies D_f(P_{XY} \| P_X \times P_Y) = \chi^2(P_{XY} \| P_X \times P_Y) := \phi^2(X, Y). \quad (14)$$

Reflexivity: If $f \in \mathcal{C}$, then $D_f(P\|Q) \geq 0$.

If f is also strictly convex at 1, then $D_f(P\|Q) = 0 \iff P = Q$.

Convexity: $D_f(P\|Q)$ is convex in (P, Q) .

Uniqueness: f and g -divergences are identical if and only if there exists a constant $c \in \mathbb{R}$ such that

$$f(t) - g(t) = c(t - 1), \quad t > 0.$$

Symmetry: let f^* be the ***-conjugate function** of $f \in \mathcal{C}$, given by

$$f^*(t) = t f\left(\frac{1}{t}\right) \tag{15}$$

for all $t > 0$. Then, $f^* \in \mathcal{C}$, and

$$D_f(P\|Q) = D_{f^*}(Q\|P). \tag{16}$$

Distance Metrics

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- **Csiszár** and Fischer considered powers of symmetrized α divergences for $\alpha \in (0, 1)$ which are distance metrics:

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- Kafka et al. (1991): If $f = f^*$ and $f(t)(1 - t^\beta)^{-\frac{1}{\beta}}$ is monotonically non-decreasing on $t \in [0, 1)$, then $D_f^\beta(P\|Q)$ is a distance metric.
- Ostreicher-Vajda (2003) and Vajda (2009) studied explicit f -divergences satisfying the above conditions by Kafka et al.
- Square-roots of f -divergences which are bounded distance metrics:
 - ▶ $d_1(P, Q) = \sqrt{\mathcal{H}^2(P\|Q)}$;
 - ▶ $d_2(P, Q) = \sqrt{D(P\|\frac{1}{2}(P+Q)) + D(Q\|\frac{1}{2}(P+Q))}$.

Data Processing Inequality (Csiszár, 1967)

Let

- $f \in \mathcal{C}$;
- $(\mathcal{X}, \mathcal{X})$ and $(\mathcal{Y}, \mathcal{Y})$ be measurable spaces;
- P and Q be probability measures on \mathcal{X} ;
- for all $x \in \mathcal{X}$, $K(\cdot|x)$ is a probability measure that is \mathcal{Y} -measurable;
- KP and KQ are prob. measures on \mathcal{Y} such that, for every $\mathcal{B} \in \mathcal{Y}$,

$$KP(\mathcal{B}) := \int_{\mathcal{X}} K(\mathcal{B}|x) dP(x), \quad KQ(\mathcal{B}) := \int_{\mathcal{X}} K(\mathcal{B}|x) dQ(x).$$

Then,

$$D_f(KP\|KQ) \leq D_f(P\|Q). \quad (17)$$

Range of Values Theorem (Vajda, 1972)

- The range of an f -divergence is given by

$$0 \leq D_f(P\|Q) \leq f(0) + f^*(0) \quad (18)$$

where

$$f^*(0) := \lim_{t \downarrow 0} f^*(t) = \lim_{u \rightarrow \infty} \frac{f(u)}{u}, \quad (19)$$

and

- ▶ $D_f(P\|Q) = 0$ if $P = Q$;
- ▶ $D_f(P\|Q) = f(0) + f^*(0)$ if $P \perp Q$ (i.e., $\text{supp}(P) \cap \text{supp}(Q) = \emptyset$);
- ▶ every value in this range is attainable by a suitable pair of (P, Q) .

Strengthened Version (Feldman and Österreicher, 1989)

$$\sup_{P \neq Q} \frac{D_f(P||Q)}{|P - Q|} = \frac{1}{2}(f(0) + f^*(0)). \quad (20)$$

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Implication

$$D_f(P||Q) \leq \frac{1}{2}(f(0) + f^*(0)) |P - Q| \quad (21)$$

if $f(0), f^*(0) < \infty$.

Local Behavior of f -divergences (Csiszár, 1967)

If $f \in \mathcal{C}$ is strictly convex at 1, then $\exists \psi_f : [0, \infty) \rightarrow [0, \infty)$ such that

- $\lim_{x \downarrow 0} \psi_f(x) = 0$;
- $|P - Q| \leq \psi_f(D_f(P\|Q))$.

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Corollary

If $f \in \mathcal{C}$ is strictly convex at 1, then

$$\lim_{n \rightarrow \infty} D_f(P_n \| Q_n) = 0 \Rightarrow \lim_{n \rightarrow \infty} |P_n - Q_n| = 0. \quad (22)$$

Special case:

Convergence to 0 in relative entropy \implies Convergence to 0 in TV distance.

Local Behavior of f -divergences (Pardo and Vajda, 2003)

Let

- $\{P_n\}$ be a sequence of probability measures on $(\mathcal{A}, \mathcal{F})$;
- the sequence $\{P_n\}$ converge to a prob. measure Q in the sense that

$$\lim_{n \rightarrow \infty} \text{ess sup} \frac{dP_n}{dQ}(Y) = 1, \quad Y \sim Q \quad (23)$$

where $P_n \ll Q$ for all sufficiently large n .

- f be convex on $(0, \infty)$, and f'' be continuous at 1.

Then,

$$\lim_{n \rightarrow \infty} \frac{D_f(P_n \| Q)}{\chi^2(P_n \| Q)} = \frac{1}{2} f''(1), \quad \lim_{n \rightarrow \infty} \chi^2(P_n \| Q) = 0. \quad (24)$$

M. Pardo and I. Vajda, "On asymptotic properties of information-theoretic divergences," *IEEE T-IT*, vol. 49, pp. 1860–1868, July 2003.

Local Behavior: Example

$$P_n \rightarrow Q \implies \lim_{n \rightarrow \infty} \frac{D(P_n \| Q)}{\chi^2(P_n \| Q)} = \frac{1}{2} \log e. \quad (25)$$

Proof

Let

$$f(t) = t \log t.$$

Then,

$$D_f(P_n \| Q) = D(P_n \| Q), \quad f''(1) = \log e.$$

Local Behavior of Relative Entropy (Csiszár, 1975)

In one of his famous papers, Csiszár proved that

$$\lim_{\lambda \downarrow 0} \frac{1}{\lambda} D(\lambda P + (1 - \lambda)Q \| Q) = 0. \quad (26)$$

Reference

I. Csiszár, “I-divergence geometry of probability distributions and minimization problems,” *Annals of Probability*, vol. 3, no. 1, pp. 146–158, 1975.

Local Behavior of f -divergences (I.S., 2018)

As a continuation to Csiszár's result (1975), we strengthened it as follows:

Let

- P and Q be prob. measures on $(\mathcal{A}, \mathcal{F})$, and suppose that

$$\text{ess sup } \frac{dP}{dQ}(Y) < \infty, \quad Y \sim Q; \quad (27)$$

- $f \in \mathcal{C}$, and its second derivative is continuous at 1.

Then,

$$\lim_{\lambda \downarrow 0} \frac{1}{\lambda^2} D_f(\lambda P + (1 - \lambda)Q \| Q) = \lim_{\lambda \downarrow 0} \frac{1}{\lambda^2} D_f(Q \| \lambda P + (1 - \lambda)Q) \quad (28)$$

$$= \frac{1}{2} f''(1) \chi^2(P \| Q). \quad (29)$$

I. Sason, "On f -Divergences: integral representations, local behavior, and inequalities," *Entropy*, May 2018.

$D(P\|Q)$ and $\chi^2(P\|Q)$ (I.S.)

Let P and Q be probability measures. Then,

$$\frac{1}{\log e} D(P\|Q) = \int_0^1 \chi^2(P \parallel (1 - \lambda)P + \lambda Q) \frac{d\lambda}{\lambda}, \quad (30)$$

$$\frac{1}{2} \chi^2(Q\|P) = \int_0^1 \chi^2((1 - \lambda)P + \lambda Q \parallel P) \frac{d\lambda}{\lambda}. \quad (31)$$

Csiszár-Kemperman-Kullback-Pinsker inequality (~ 1967)

$$D(P\|Q) \geq \frac{1}{2} |P - Q|^2 \log e.$$

A Simple Proof (I.S.)

- By invoking the Cauchy-Schwartz, it readily follows that

$$\chi^2(P\|Q) \geq |P - Q|^2. \quad (32)$$

- Using (32), we get

$$\begin{aligned} \frac{1}{\log e} D(P\|Q) &= \int_0^1 \chi^2(P \parallel (1 - \lambda)P + \lambda Q) \frac{d\lambda}{\lambda} \\ &\geq \int_0^1 \underbrace{|P - ((1 - \lambda)P + \lambda Q)|^2}_{=\lambda^2 |P - Q|^2} \frac{d\lambda}{\lambda} \\ &= |P - Q|^2 \int_0^1 \lambda d\lambda \\ &= \frac{1}{2} |P - Q|^2. \end{aligned}$$

Other Simple Proofs

Apart of the Csiszár-Kemperman-Kullback-Pinsker inequality, the identity

$$\frac{1}{\log e} D(P\|Q) = \int_0^1 \chi^2(P\|(1-\lambda)P + \lambda Q) \frac{d\lambda}{\lambda}$$

enables us to prove several (new and old) f -divergence inequalities:

$$D(P\|Q) \leq \frac{1}{3} \chi^2(P\|Q) + \frac{1}{6} \chi^2(Q\|P), \quad (33)$$

$$D(P\|Q) \leq \frac{1}{2} \chi^2(P\|Q) + \frac{1}{4} |P - Q|. \quad (34)$$

Pinsker-Type Inequalities for f -divergences

Theorem (Csiszár, 1966)

Let $f \in \mathcal{C}$ be twice differentiable, and let $r_0 \in (0, 1)$ and $a > 0$ satisfy

$$f''(u) \geq a > 0, \quad \forall u \in (1 - r_0, 1 + r_0). \quad (35)$$

Let $\delta \leq r_0^2$. Then,

$$D_f(P\|Q) \leq \delta \implies |P - Q| \leq c\sqrt{\delta}, \quad (c := \frac{2}{a} + 1). \quad (36)$$

$f(t) = t \ln t$ for $t > 0$, and $r_0 = \frac{1}{2}$, $a = \frac{2}{3} \implies |P - Q| \leq 4\sqrt{D(P\|Q)}$ nats.

1-to-1 correspondence $\mathcal{I}_\alpha \leftrightarrow D_\alpha \implies$ extendable to the Rényi divergence.

1. Csiszár, "A note on Jensen's inequality," *Studia Scient. Math. Hungarica*, 1966.

Csiszár-Kemperman-Kullback-Pinsker Inequality

$$\inf_{P \neq Q} \frac{D(P||Q)}{|P - Q|^2} = \frac{1}{2} \log e \quad \Longrightarrow \quad D(P||Q) \geq \frac{1}{2} |P - Q|^2 \log e. \quad (37)$$

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Question

Is there a [reverse Pinsker inequality](#) which provides an upper bound on the relative entropy as a function of the TV distance ?

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No, for every $\varepsilon > 0$, $\exists (P, Q)$ s.t. $|P - Q| \leq \varepsilon$, $D(P\|Q) = \infty$. ☹

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However, we can obtain a reverse Pinsker inequality when the relative information is **bounded**. ☺

Reverse Pinsker Inequality: Finite Alphabet (Csiszár & Talata, 2006)

If \mathcal{A} is a finite set, P and Q are probability measures defined on \mathcal{A} , and $Q_{\min} := \min_{x \in \mathcal{A}} Q(x) > 0$, then

$$D(P\|Q) \leq \frac{\log e}{Q_{\min}} \cdot |P - Q|^2. \quad (38)$$

Recent Applications of (38)

- I. Csiszár and Z. Talata, "Context tree estimation for not necessarily finite memory processes, via BIC and MDL," *IEEE T-IT*, Mar. 2006.
- V. Kostina and S. Verdú, "Channels with cost constraints: strong converse and dispersion," *IEEE T-IT*, May 2015.
- K. Marton, *Distance-divergence inequalities: rate of decrease of divergence (from stationary distribution) for Gibbs samplers*, ISIT 2013 Shannon lecture, July 2013.
- M. Tomamichel and V. Y. F. Tan, "A tight upper bound for the third-order asymptotics for most discrete memoryless channels," *IEEE T-IT*, Nov. 2013.

β_1 and β_2

Given a pair of probability measures (P, Q) on the same measurable space, denote $\beta_1, \beta_2 \in [0, 1]$ by

$$\beta_1 = \exp(-\text{ess sup } \iota_{P\|Q}(Y)), \quad (39)$$

$$\beta_2 = \text{ess inf } \exp(\iota_{P\|Q}(Y)) \quad (40)$$

with $Y \sim Q$.

A Reverse Pinsker Inequality (I.S. and S. Verdú, 2016)

If $\beta_1 \in (0, 1)$ and $\beta_2 \in [0, 1)$, then,

$$D(P\|Q) \leq \frac{1}{2} (\varphi(\beta_1^{-1}) - \varphi(\beta_2)) |P - Q| \quad (41)$$

where $\varphi: [0, \infty) \mapsto [0, \infty)$ is given by

$$\varphi(t) = \begin{cases} 0 & t = 0 \\ \frac{t \log t}{t - 1} & t \in (0, 1) \cup (1, \infty) \\ \log e & t = 1. \end{cases} \quad (42)$$

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Generalized to Rényi divergences of order $\alpha \in (0, \infty)$.

I.S. and S. Verdú, "f-divergence inequalities," *IEEE T-IT*, Nov. 2016.

A Reverse Pinsker Inequality (Binette, 2018)

For fixed $\delta \in [0, 2]$, $\beta_1, \beta_2 \in [0, 1]$, let $\mathcal{D}(\delta, \beta_1, \beta_2)$ denote the set of all probability measures P and Q with

$$|P - Q| = \delta, \quad (43)$$

$$\beta_1 = \exp(-\text{ess sup } \iota_{P\|Q}(Y)), \quad (44)$$

$$\beta_2 = \text{ess inf } \exp(\iota_{P\|Q}(Y)) \quad (45)$$

where $Y \sim Q$. Then,

$$\max_{(P,Q) \in \mathcal{D}(\delta, \beta_1, \beta_2)} D_f(P\|Q) = \frac{1}{2} \delta \left(\frac{f(\beta_1^{-1})}{\beta_1^{-1} - 1} + \frac{f(\beta_2)}{1 - \beta_2} \right) \quad (46)$$

and the maximum is attained by P and Q defined on a set of size 3. Specialized to the result for relative entropy with a similar proof's concept.

O. Binette, "Note on reverse Pinsker inequalities," May 15, 2018. [Online]. Available at <https://arxiv.org/abs/1805.05135>.

f -Informativity (Csiszár 1972)

I. Csiszár, "A class of measures of informativity of observation channels," *Periodica Mathematicarum Hungarica*, vol. 2, pp. 191–213, Mar. 1972.

f -informativity measures generalize Shannon's mutual information, and Gallager's function E_0 in the random coding error exponent.

f-Informativity (Cont.)

Let

- $f \in \mathcal{C}$;
- $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ be a family of probability measures defined on \mathcal{X} ;
- w be a probability measure defined on Θ .

The f -informativity, $I_f(w, \mathcal{P})$, is defined as

$$I_f(w, \mathcal{P}) := \inf_Q \int_{\Theta} D_f(P_\theta \| Q) \, dw(\theta) \quad (47)$$

where the infimum is taken over all probability measures Q on \mathcal{X} .

Special Case of f -informativity: Mutual Information

Let $f(t) = t \log t$ for $t > 0$, then $D_f(\cdot \| \cdot) = D(\cdot \| \cdot)$, and

$$\begin{aligned} I_f(w, \mathcal{P}) &:= \inf_Q \int_{\Theta} D(P_{\theta} \| Q) \, dw(\theta) \\ &= \int_{\Theta} D(P_{\theta} \| Q^*) \, dw(\theta) \end{aligned} \quad (48)$$

with

$$Q^*(x) := \int_{\Theta} P_{\theta}(x) \, dw(\theta), \quad \forall x \in \mathcal{X}. \quad (49)$$

This follows from the identity by Topsøe (Stud. Sci. Math. Hung., 1967):

$$\int_{\Theta} D(P_{\theta} \| Q) \, dw(\theta) = \int_{\Theta} D(P_{\theta} \| Q^*) \, dw(\theta) + D(Q^* \| Q). \quad (50)$$

Hence, the f -informativity is specialized to the mutual information:

$$I_f(w, \mathcal{P}) = I(X; \theta). \quad (51)$$

Properties

In the Bayesian case, f -informativities share several useful properties of the mutual information, such as the [data processing inequality](#).

Absolute f -Informativity (f -Radius)

$$\rho_f(\mathcal{P}) = \inf_Q \sup_{\theta \in \Theta} D_f(P_\theta \| Q). \quad (52)$$

Hence,

$$0 \leq I_f(w, \mathcal{P}) \leq \rho_f(\mathcal{P}), \quad (53)$$

so, the non-negative f -informativity is upper bounded by the f -radius.

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so, the non-negative f -informativity is upper bounded by the f -radius.

- For observation channels without prior probabilities, f -informativities have the geometric interpretation of a radius.
- In view of the redundancy-capacity theorem, the f -radius is a generalization of the channel capacity (let $f(t) = t \log t$ for $t > 0$).

Parameter Estimation: Basic Model

- The **estimand** is an unknown parameter θ ;
- Θ is the **parameter space** for θ ;
- \mathcal{X} is the **sample space** for the observed data X ;
- $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ is the **model** for the data X conditioned on θ ;
- \mathcal{A} is the **action space** for the estimation $\hat{\theta}$, based on the data X ;
- $L: \Theta \times \mathcal{A} \mapsto [0, \infty)$ is a **loss function** for estimating θ by $\hat{\theta}$.

Estimator

Let $T: \mathcal{X} \mapsto \mathcal{A}$ be an arbitrary mapping where $\hat{\theta} = T(x)$ for $x \in \mathcal{X}$.

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Risks

- **Minimax risk:** expected loss for the best estimator & worst prior

$$R_{\text{minimax}}(L; \Theta) := \inf_{T: \mathcal{X} \mapsto \mathcal{A}} \sup_{\theta \in \Theta} \mathbb{E}[L(\theta, T(X))] \quad (54)$$

where the expectation is taken over $X \sim P_{\theta}$.

- **Bayes risk:** expected loss for the best estimator with a prior w on Θ

$$R_{\text{Bayes}}(w, L; \Theta) := \inf_{T: \mathcal{X} \mapsto \mathcal{A}} \int_{\Theta} \mathbb{E}[L(\theta, T(X))] dw(\theta). \quad (55)$$

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$$\implies R_{\text{minimax}}(L; \Theta) \geq R_{\text{Bayes}}(w, L; \Theta) \quad \forall \text{ prior } w.$$

Bayes Risk

- If the prior distribution w is known, then the Bayes estimator attains the Bayes risk; 😊
- In general, however, the Bayes estimator is computationally hard to evaluate \implies Bayes risk has, often, no closed-form expression. 😞

Bayes Risk Lower Bound

A lower bound on the Bayes risk

- characterizes the fundamental limit of any estimator given the prior knowledge;
- serves as a lower bound on the minimax risk (for the worst prior).

Bayes Risk Lower Bounds (Cont.)

Approach

- Derivation of Bayes risk lower bounds relies heavily on the data processing inequality for f -divergences.
- First derived for 0 – 1 loss functions, and then extended to an arbitrary non-negative loss function.

Reference

X. Chen, A. Guntuboyina, and Y. Zhang, “On Bayes risk lower bounds,” *Journal of Machine Learning Research*, vol. 17, pp. 1–58, Dec. 2016.

Notation

Let $f \in \mathcal{C}$, and define the function $\phi_f: [0, 1]^2 \mapsto \mathbb{R}$ as follows:

$$\phi_f(a, b) := \begin{cases} bf\left(\frac{a}{b}\right) + (1-b)f\left(\frac{1-a}{1-b}\right), & (a, b) \in [0, 1] \times (0, 1) \\ af^*(0) + f(1-a), & (a, b) \in [0, 1] \times \{0\} \\ f(a) + (1-a)f^*(0), & (a, b) \in [0, 1] \times \{1\}. \end{cases}$$

Bayes Risk Lower Bounds for Arbitrary Non-Negative Loss Functions

For arbitrary

- $f \in \mathcal{C}$;
- upper bound \bar{I}_f on the f -informativity $I_f(w, \mathcal{P})$,

let $u_f: [0, \infty) \mapsto [\frac{1}{2}, 1]$ be the monotonically non-decreasing function:

$$u_f(x) := \inf \left\{ \frac{1}{2} \leq b \leq 1 : \phi_f\left(\frac{1}{2}, b\right) > x \right\}, \quad x \geq 0 \quad (56)$$

and if $\phi_f\left(\frac{1}{2}, b\right) \leq x$ for every $b \in [\frac{1}{2}, 1]$, then $u_f(x) := 1$. Then,

$$R_{\text{Bayes}}(w, L; \Theta) \geq \frac{1}{2} \sup \left\{ t > 0 : \sup_{a \in \mathcal{A}} w(B_t(a, L)) < 1 - u_f(\bar{I}_f) \right\} \quad (57)$$

where, for $a \in \mathcal{A}$ and $t > 0$,

$$B_t(a, L) := \left\{ \theta \in \Theta : L(\theta, a) < t \right\}. \quad (58)$$

Bayes Risk Lower Bounds (Cont., Chen *et al.*, 2016)

Specialization to specific f -divergences yields the following lower bounds:

- Relative entropy ($f(t) = t \log t$ for $t > 0$):

$$R_{\text{Bayes}}(w, L; \Theta) \geq \frac{1}{2} \sup \left\{ t > 0 : \sup_{a \in \mathcal{A}} w(B_t(a, L)) < \frac{1}{2} - \frac{1}{2} \sqrt{1 - \exp(-2\bar{I}_{\text{KL}})} \right\} \quad (59)$$

- χ^2 divergence ($f(t) = t^2 - 1$ for $t > 0$):

$$R_{\text{Bayes}}(w, L; \Theta) \geq \frac{1}{2} \sup \left\{ t > 0 : \sup_{a \in \mathcal{A}} w(B_t(a, L)) < \frac{1}{2} - \frac{1}{2} \sqrt{\frac{\bar{I}_{\chi^2}}{1 + \bar{I}_{\chi^2}}} \right\}. \quad (60)$$

f -Divergences and f -Informativities: Theory and Applications

- I-divergence (relative entropy), and generalization to f -divergences;

f -Divergences and f -Informativities: Theory and Applications

- I-divergence (relative entropy), and generalization to f -divergences;
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f -Divergences and f -Informativities: Theory and Applications

- **I**-divergence (relative entropy), and generalization to f -divergences;
- **M**utual information, and generalization by means of f -informativities;
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Best wishes !

Legjobbkat kívánom !